Contribution of machine learning to the analysis of grade repetition in Spain: A study based on PISA data Contribución del machine learning al análisis de la repetición

escolar en España: un estudio con datos PISA

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Abstract:

Introduction: The rate of grade repetition is excessively high in Spain despite being a controversial measure. In order to obtain evidence to contribute to reducing it in compulsory education, the present work is an in-depth study of the PISA 2018 context indices that are most closely linked to this phenomenon. Method: With the sample of Spanish students (n =35943), we used an automatic machine learning method to select and order the predictors, and multilevel logistic regression (students and centres) to quantify the contribution of each one. Results: For each educational stage we obtained the 30 most significant contextual variables, which explain 65.5% of the grade repetition variance in primary education and almost 55.7% in secondary education. Conclusions: The main indicators are principally at student level, which suggests the suitability of psychoeducational interventions based more on individual support than general policies. This gives rise to potentially more efficient and equitable measures than grade repetition, aimed at, for example, the management of learning time or academic/professional guidance, and predictors with specific differential significance at each stage. Methodologically, the study contributes to improving the specification of predictive models. **Revista Española de Pedagogía** year 82, n. 289, September-December 2024, 539-562

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Keywords: PISA, grade repetition, machine learning, contextual variables, multilevel logistic regression, compulsory education.

Resumen:

Introducción: La repetición de curso tiene una excesiva incidencia en España a pesar de ser una medida controvertida. A fin de obtener evidencias que contribuvan a su reducción en la educación obligatoria, el presente trabajo profundiza en el estudio de los índices de contexto de PISA 2018 más vinculados con dicho fenómeno. Método: Con la muestra de estudiantes españoles (n = 35943), se utiliza un método de aprendizaje automático para seleccionar y ordenar los predictores, y una regresión logística multinivel (estudiantes y centros) para cuantificar la contribución de cada uno. Resultados: Se obtienen las 30 variables de contexto más relevantes en cada etapa educativa, que explican el 65.5% de la varianza de la repetición en primaria y casi el 55.7% en secundaria. Conclusiones: Los principales indicadores son sobre todo del nivel de estudiantes, lo que sugiere la idoneidad de intervenciones psicoeducativas basadas en el apoyo individualizado más que en políticas generalizadas. De ahí emergen medidas potencialmente más eficientes y equitativas que la repetición, centradas, por ejemplo, en la gestión del tiempo de aprendizaje o en la orientación académico-profesional, así como predictores con importancia específica diferencial en cada etapa. En el ámbito metodológico, el estudio hace una aportación a la mejora de la especificación de los modelos predictivos.

Palabras clave: PISA, repetición escolar, aprendizaje automático, variables de contexto, regresión logística multinivel, educación obligatoria.

1. Introduction

As is generally known, grade repetition is a strategy that consists of retaining the student at a level of a specific educational stage when they have not proved that they have acquired the minimum levels of knowledge (Jimerson & Ferguson, 2007; López & García, 2020). Nevertheless, it is a controversial measure, that some authors consider to be positive because it helps lower-performing students to improve their skills and gain maturity (Battisttin & Schizzerotto, 2019; Ikeda & García, 2014; Valbuena et al., 2021) and that others see as a merely disciplinary measure (Schwerdt et al., 2017). Furthermore, although this issue has been the subject of synthesis studies, systematic reviews and meta-analysis covering practically the entire twentieth century (Allen et al., 2009; Jimerson 2001), there are doubts as to its effectiveness. Moreover, it even seems to be detrimental, as it is associated with a decrease in student performance (Asensio et al., 2018; Hattie, 2017; López et al., 2018, Nieto-Isidro & Martínez-Abad, 2023) and is linked to negative aspects, such as frustration (López-Rupérez et al., 2021), feelings of discrimination



(Van Canegem et al., 2021) or a lowering of self-confidence, academic self-concept or motivation, thus affecting academic expectations (González-Nuevo et al., 2023; Mathys et al., 2019; Peixoto et al., 2016).

The corresponding regulations in Spain take this situation into account and the Organic Law 1/1990 of 3 October, on the General Organisation of the Educational System (LOGSE) established that repetition is a measure that focuses on exceptional diversity. Although, in fact, its use is not so exceptional. As reported by the Ministerio de Educación y Formación Profesional [Ministry of Education and Vocational Training] (2022), in the first stage of secondary education the repetition rate is 8.7% and in primary education 2.4% of students have repeated a year, a percentage that is only surpassed in the European Union by Portugal (3.6%), the Slovak Republic (2.9%)and Austria (2.9%). Furthermore, based on data from the Programme for International Student Assessment (PISA) for 2009, 2015 and 2018, López and García (2020), who compare the percentage of 15-year-old students who have repeated at least once in the European Union, conclude that Spain, despite having considerably reduced the rate between 2009 and 2018, in the latter year still has a rate of 28.71%. That puts the country in third place, only behind Luxembourg (32.02%) and Belgium (29.06%), when the average for the Organisation for Economic Co-operation and Development (OECD) stands at 11.4% and for Finland it is 3.2%.

In view of the above, further research into this phenomenon is necessary, with the ultimate aim of providing information that makes it possible to act based on the evidence, not only from a political perspective, but also through educational practice, thereby contributing to data optimisation so that Spain can get closer, at least, to the OECD average.

To this effect, the present study takes the line of a search for predictors by means of secondary exploitation of the PISA data, for which there is already research based on previous PISA publications, showing that grade repetition has many causes (Arroyo et al., 2019; Carabaña, 2013; Cordero et al., 2014).

In line with the classification of background questionnaires used by PISA (OECD, 2019), shown in Figure 1 below, there is a brief summary of the state of affairs regarding the most significant variables of the three constructs.

Regarding precedents, there is evidence that, above all, the students who repeat grades usually come from underprivileged socio-economic backgrounds (López & García, 2020). However, immigration background shows inconclusive results, as Cordero (2014), García-Pérez et al. (2014) o Warren et al. (2014) indicate that first-generation immigrants (born abroad) have a greater probability of repeating a grade than natives, while, for example, in the study by Choi et al. (2018) this variable does not prove to be significant. Revista Española de Pedagogía year 82, n. 289, September-December 2024, 539-562

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FIGURE 1. Modular structure of contextual variables in PISA 2018.

	Student background constructs	d constructs		Schooling constructs		Non-cognitive and metacognitive constructs
			$TE\ell$	TEACHING AND LEARNING	NG	
y literacy		5. Out-of scool	1. Teacher qualifica- tions and knowledge	2. Teaching practices for science	11. Learning time and curriculum	4. Reading related outcomes: attitudes, motivation and strategies
Buip		scientific experience		SCHOOL POLICIES		
Res				3. School-level learn- ing environment for reading		
egories	6. Socio-economic status	8. 	13. Implicación de los padres	12. School climate: interpersonal relationships, trust and expectations	14. School context and resources	9. Dispositional and school-focused variables
l cat	of student and family	pathways in		GOVERNANCE		
General	7. Migration and culture	early child- hood		16. Student assessment, institutional assess- ment and account- ability	15. Allocation, selection and choice	10. Dispositions for global competence

Secondly, among the school-related constructs, studies have mainly been conducted on learning time, school climate or the curriculum, which are associated with grade repetition (OECD, 2019). In this way, Asensio et al. (2018) and López et al. (2018) demonstrate the importance of the number of classes per week or the duration of classes for grade repetition, while Seabra and Ferrão (2016) find a significant relationship between grade repetition and disciplinary problems in schools, with the disciplinary climate having a greater impact than the socio-economic level. On the other hand, Arroyo et al. (2019) link repetition with curriculum-related variables, such as having studied science the previous year or having chosen an optional subject in this field.

In the group of non-cognitive and metacognitive constructs (hereinafter NC+MCC), motivation, self-confidence, future academic expectations and a sense of belonging to the group are highly significant variables (Fernández-Lasarte et al., 2019; Hornstra et al., 2017; Van Canegem et al., 2021), so that positive development of these aspects minimises the effect of grade repetition (Marsh et al., 2018). Likewise, academic self-concept and focusing on goals provide a high level of predictive capacity regarding the phenomenon under study (Ferla et al., 2009; Rodríguez-Rodríguez, 2022).

Despite the amount of existing research, one limitation that makes it difficult to define a more accurate picture lies in the fact that predictors are usually introduced into models in a random way, perhaps due to the lack of theoretical references that, in education, would produce modelling that aims to minimise specification error. Although the main assumption of multilevel models (which are the most widely used in secondary exploitation of PISA data, due to their isomorphism with the reality they aim to model) is that they are well defined, that is to say, that they meet the condition of "not excluding major predictors from the model" (Gaviria & Castro, 2005, p. 86).

In view of this, it is interesting to explore predictive models in which statistical error reduction is matched by the attempt to minimise specification error as well. The present research takes this line, as do others that use data mining, based on machine learning (Urbina & De la Calleja, 2017), both in international studies (Gamazo & Martínez-Abad, 2020; Kilıç et al., 2017; Kiray et al., 2015; Liu & Ruiz, 2008; Martínez-Abad & Chaparro, 2017; Martínez-Abad et al., 2020), and those focusing on Spanish samples (Arroyo et al., 2024a; Arroyo et al., 2024b; Asensio et al., 2018).

Another limitation is that the previous research does not tend to distinguish between students from the two compulsory stages, with most studies concentrating on secondary school students, which prevents the appearance of differentially significant variables in primary education.

In order to contribute to overcoming the aforementioned limitations, the overall objective pursued by this research is to obtain pertinent information for modelling that is a closer fit for grade repetition in the two compulsory stages in Spain, based on the PISA 2018 databases (Ministerio de Educación y Formación Profesional, 2019), which provide a vast amount of information about contextual variables (OECD, 2019).

The specific goals are:

- To identify the most accurate machine learning algorithm [individual decision trees (CART and C5.0), random forest and stochastic gradient boosting (GBM)] and, with it, select the variables that are most closely linked to grade repetition in the two stages and sort them according to their differential significance in each stage.
- To determine the relative contribution of each factor (student background, school-related and NC+MCC constructs) to the probability of a student repeating a grade in primary and secondary education taking into account the two-level structure of the data (students and schools).

2. Method

A secondary analysis was conducted of the PISA 2018 international assessment data, using a non-experimental, quantitative, cross-disciplinary, correlational design, in accordance with the PISA manual on technical guidelines (OECD, 2019).

2.1. Participants

Using two-stage sampling, the sample of Spanish students who participated in

PISA 2018 is 35943, which represents coverage of 92% of the national population of 15-year-olds (Ministerio de Educación y Formación Profesional, 2019), composed of 49.96% women and with an average age of 15.84 years old (SD =0.29). By type of school, 65% are students who are enrolled in public centres, 28% in state-contracted, privately-owned centres and 7% in private schools. The full information can be found in the second annex of the document published by the OECD (2020).

2.2. Instruments

All the variables have been measured with the PISA 2018 background questionnaires for students and principals (OECD, 2019). The dependent variables are grade repetition in primary and secondary education, with almost double the rate in the latter stage as reported by the participants themselves (Table1).

In terms of the independent variables, 85 were introduced in the initial stage (see Tables S1 and S2 at https:// acortar.link/hfDnvO). Of this set of predictors, 83 are composite indices (WLE and SUM), of which 67 indices were constructed based on student responses (level 1) and 17 based on the school principals' responses (level 2). The two remaining variables are simple items that have been included in the model due to the large amount of previously existing evidence: student gender (level 1) and percentage of 15-year-old students from socio-economically disadvantaged homes (level 2).

Have you	Have you ever repeated a grade in primary education?						
	n	%	n (weighted)	% (weighted)			
No	30078	89.927	340287	88.606			
Yes, once or more	3369	10.073	43756	11.394			
Total valid responses	33447						
Have you	Have you ever repeated a grade in secondary education?						
	n	%	n (weighted)	% (weighted)			
No	27902	80.354	310497	77.313			
Yes, once or more	6822	19.646	91114	22.687			
Total valid responses	34724						

TABLE 1. Grade repetition in compulsory education in Spain.

Source: author's own compilation based on the PISA (2018) data for Spain.

Subsequently, the multilevel models included only those that, using the more accurate machine learning algorithm, show a greater connection with grade repetition in each of the stages studied.

2.3. Procedure

A detailed summary of the procedure for applying the questionnaires can be found in the Spanish report (Ministerio de Educación y Formación Profesional, 2019).

2.4. Data analysis

There were two separate stages, so that first an exploratory, descriptive study was conducted (Rosenthal & Rosnow, 2008) using algorithms supervised by machine learning and, once the most significant variables had been selected and sorted, several predictive models were tested taking into account the hierarchical data structure (level 1: student and level 2: school).

Prior to the machine learning algorithm, the data were pre-processed, as recommended by Raschka (2015), by processing the missing values with the technique of multiple imputation by chained equations (Rivero, 2011) and proceeding to random division of the cases into two sets: one for training (70%), with which to adjust the model, and the other for validation (30%), with which to evaluate the model by means of the area under the curve (AUC), as these are two dependent variables with unbalanced levels. Centring and scaling were also performed on the continuous predictors in order to avoid overrepresentation of the variables with a greater magnitude, and no predictor with zero or near-zero variance was obtained.

To meet the first objective, we compared the performance of four algorithms [individual decision trees (CART and C5.0), random forest and stochastic

gradient boosting] using the AUC, with the grid-search method based on 10-fold cross-validation, jointly optimising the most significant hyperparameters. Using the sorting method of relative importance, Stochastic Gradient Boosting proved to be the most accurate algorithm, in accordance with the Friedman test applied to the training sample. For the sake of parsimony, given that one of the limitations of this algorithm is that it does not suggest a fixed cut-off point regarding the selection of an optimal number of predictors (Sarkar et al., 2018), in line with Gorostiaga and Rojo-Álvarez (2016), we decided to examine the following sets of variables: 15, 20, 25, 30 and 35. In this way, greater precision was provided for the selection of the optimal set, which proved to be, in both primary and secondary education, the one containing 30 of all the indices initially included.

To achieve the second objective, we used the technique of multilevel binary logistic regression, as the random variance of the null model proved to be statistically significant and the intraclass correlation coefficient (ICC) was greater than 10% (Lee, 2000). In considering the assumption of multicollinearity, we eliminated *home possessions* (HOMEPOS), *mother's education* (MISCED) and *parents' highest occupational status* (HISEI), as they are implicit in the *index of economic, social and cultural status* (ESCS); and the predictor *family wealth*, due to its strong correlation with this index.



Following this, as required by multilevel analysis, we removed the schools with fewer than 20 subjects, thereby excluding 1532 subjects belonging to 133 centres, so the final sample consisted of 34411 students.

This generated a total of five models: the null model, without predictors; model 1, composed of the student background variables; model 2, comprising the school-related variables; model 3, with the NC+MCC variables; and model 4, containing all the predictors. Thus, it was possible to analyse the contribution of the school-related and NC+MCC factors once the student background had been checked. The variables were introduced into the model in the order obtained from the machine learning algorithm (Arroyo et al., 2019; Constante-Amores et al., 2021). Finally, to check the fit of the models, we used the AIC and BIC indices and Deviance, a statistical reduction that allows nested models to be compared by estimating the percentage of reduced variance (R2), and we calculated their significance (Cameron & Windmejier, 1997). The interpretation of the results is based on the percentage of explained variance (PEV), which quantifies the explanatory capacity of the final model, and on the odds ratios, in which values greater than 1 indicate that the variable increases the probability of repeating a grade, acting as a risk factor, and values below 1 suggest that this probability decreases, acting as a protective factor.

In terms of software, we used the version 4.4.0 R packages caret, stochastic gradient boosting and lme4 to compare the machine learning algorithms, select the most important variables and conduct multilevel binary logistic regression, respectively.

3. Results

Of the algorithms tested, the most accurate were stochastic gradient boosting and random forest, for both primary education (Figure 2) and secondary education (Figure

3). With the Friedman test, of note is stochastic gradient boosting, which we eventually used to select and hierarchise the most significant variables associated with grade repetition in compulsory education.

FIGURE 2. Comparison of machine learning algorithms for primary education.



FIGURE 3. Comparison of machine learning algorithms for secondary education.



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Table 2 shows the empirical results produced by stochastic gradient boosting, in which, out of the 30 most significant variables in primary education, 12 relate to background, 9 are associated with the school and 9 are NC+MCC. All of them correspond to level 1, except *the percentage of students from socio-economically* *disadvantaged homes*, which is from level 2, ranked 23rd. In this educational stage, the most significant predictors are the following: first, the *index of economic, so-cial and cultural status* (ESCS); secondly, the *student's expected occupational status* (BSMJ); and thirdly, *learning time for mathematics* (MMINS).

TABLE 2. Most significant variables associated with grade repetition in primary education.

	Label	Order	Variables
	ESCS	1^{st}	Index of economic, social and cultural status
	HOMEPOS	$4^{ m th}$	Home possessions
	SCCHANGE	5^{th}	Number of school changes
F	HISEI	8^{th}	Parents' highest occupational status
ouno	MISCED	9^{th}	Mother's education
Student background	IMMIG	$12^{ m th}$	Immigration background
t bac	ICTRES	16^{th}	Home ICT resources
dent	HEDRES	$17^{ m th}$	Home educational resources
Stu	DURECEC	20^{th}	Duration of early childhood education and care
	SC048Q03NA	$23^{ m rd}$	Percentage of students from socio-economically disad- vantaged homes
	BMMJ1	$27^{ m th}$	Mother's ISEI
	WEALTH	30^{th}	Family wealth
	MMINS	3^{rd}	Learning time for mathematics (minutes per week)
	LMINS	6^{th}	Learning time for language (minutes per week)
0	INFOCAR	10^{th}	Information about careers
cho	SMINS	$11^{ m th}$	Learning time for sciences (minutes per week)
tos	DIRINS	13^{th}	Teacher-directed instruction
ted	TMINS	$14^{ m th}$	Learning time in total (minutes per week)
Related to school	PERFEED	18^{th}	Perceived teacher feedback
	DISCLIMA	28^{th}	Disciplinary climate in language classrooms (core classes in this PISA publication)
	BEINGBULLIED	29^{th}	Experience of being bullied

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e	BSMJ	$2^{ m nd}$	Student's expected occupational status
metacognitive	PISADIFF	$7^{ m th}$	Perception of difficulty of the PISA test
acog	METASUM	15^{th}	Metacognition: summarising
meta	STUBMI	19^{th}	Student's body mass index
and	GCSELFEFF	21^{st}	Self-efficacy regarding global competence
1	BELONG	$22^{ m nd}$	Subjective well-being: sense of belonging at school
Non-cognitive	SCREADDIFF	24^{th}	Self-concept of reading: perception of difficulty
on-co	GCAWARE	25^{th}	Student awareness of global issues
ž	COGFLEX	26^{th}	Student's cognitive flexibility/adaptability

Table 3 presents the variables that emerge as being the most important, with Stochastic Gradient Boosting, in secondary education: 8 relating to student background, 14 associated with school and 8 non-cognitive and metacognitive related to the student. In this stage, there are 29 from level 1 and only one from level 2: *student behaviour hindering learning* (ranked 11th). The most significant predictors here are revealed as: the *student's expected occupational status* (BSMJ), *learning time for sciences* (SMINS) and the index of economic, social and cultural status (ESCS). The descriptive analyses of the 30 most important variables shown in Table 2 and Table 3 can be found in the following link: https://acortar.link/hfDnvO (Tables S3 and S4).

TABLE 3. Most significant variables associated with grade repetition in secondary education.

Area	Label	Order	Variables
	ESCS	$3^{\rm rd}$	Index of economic, social and cultural status
F	CHANGE	$5^{ m th}$	Number of changes in the educational biography
Student background	HOMEPOS	9^{th}	Home possessions
ickgr	BMMJ1	10^{th}	Mother's ISEI
nt ba	HISEI	$12^{ m th}$	Parents' highest occupational status
inder	HEDRES	20^{th}	Home educational resources
S	ICTRES	$21^{ m st}$	Home ICT resources
	DURECEC	$23^{ m rd}$	Duration of early childhood education and care

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	SMINS	$2^{ m nd}$	Learning time for sciences (minutes per week)
	LMINS	4^{th}	Learning time for Language (minutes per week)
	MMINS	6^{th}	Learning time for Mathematics (minutes per week)
	INFOCAR	$7^{ m th}$	Information about careers
	STUBEHA	$11^{ m th}$	Student behaviour hindering learning
ol	HOMESCH	13^{th}	Use of ICT at home for school work
scho	TMINS	$15^{ m th}$	Learning time in total (minutes per week)
d to	ICTOUTSIDE	$17^{ m th}$	Subject-related use of ICT outside of school
Related to school	ENTUSE	18^{th}	Use of ICT outside of school for leisure
L R	INFOJOB2	19^{th}	Information about the labour market provided outside of school
	PERFEED	$22^{ m nd}$	Perceived teacher feedback
	INFOJOB1	24^{th}	Information about the labour market provided by the school
	DISCRIM	$27^{ m th}$	Discriminatory school climate
	DIRINS	29^{th}	Teacher-directed instruction
ve	BSMJ	1^{st}	Student's expected occupational status
gniti	MASTGOAL	8^{th}	Learning goals
taco	PISADIFF	$14^{ m th}$	Perception of difficulty of the PISA test
d me	GCSELFEFF	16^{th}	Self-efficacy regarding global competence
Non-cognitive and metacognitive	RESILIENCE	25^{th}	Student resilience
gnitiv	METASUM	26^{th}	Metacognition: summarising
300-U	STUBMI	28^{th}	Student's body mass index
Noi	GCAWARE	30^{th}	Student awareness of global issues
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To achieve the second objective, once we had identified and sorted the variables with the strongest relationship to grade repetition, the next step was to define the predictive model, for which we dichotomised the categorical independent variables as indicated by Pardo and Ruiz (2013) (Table 4). For the predictor DURECEC, the OECD (2019) guidelines were followed.

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Label	Categorical independent variable	Recoded values
SCCHANGE	Number of school changes	0 = No change
CHANGE	Number of changes in the educational biography	1 = Once or more changes
IMMIG	Immigration background	0 = Immigrant 1 = Native
DURECEC	Duration of early childhood education and care	0 = Three or under 1 = Four or more

TABLE 4. Recoding of independent variables for the multilevel logistic regression model.

Table 5 summarises the regression models for grade repetition in primary education, with N1 or N2 after the variable name to identify the level to which it belongs. In model 1, all the student backgrounds are statistically significant, which explains the 62.380% variability in grade repetition. According to the odds ratio, the predictor number of school changes (SCCHANGE) presents the greatest probability of repetition. It should be noted that the *percentage* of students from socioeconomically disadvantaged homes (SC048Q03NA) is the only significant variable from level 2, and mother's ISEI (BMMJ1) has an odds ratio that is close to 1. Model 2, composed of the school-related variables, explains 19% of the criterion variance and, although they are all significant, four of them have an odds ratio of 1, with experience of being bullied (BEINGBULLIED) and teacher-directed instruction (DIRINS) being the main risk factors. Model 3 explains about 25% of the variance, containing only NC+MCC variables, of which student awareness of global issues (GCAWARE) is not significant and has an odds ratio that is very close to the body mass index (STUBMI), with perception of difficulty of the PISA test (PISADIFF) being the variable with the greatest effect. The last model (M4) explains 61% of the variance in grade repetition in primary education; there is no significance attached to the effects of learning time for sciences (minutes per week) (SMINS), or the body mass index (STUBMI), and furthermore, there are eight variables with an odds ratio that is very close to 1; the variables with the most effect are number of school changes (SCCHANGE), teacher-directed instruction (DIRINS) and perceived teacher feedback (PERFEED).

Lastly, as Table 5 shows, regarding the model fit, model 4 has the lowest AIC and BIC score, with a significant reduction in the variance, equivalent to an R^2 of 9% compared to model 1. This is also significant in relation to the null model, with a reduction in the variance of model 1 (8%), model 2 (5%) and model 3 (9%).

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TABLE 5. Odds ratio and standard errors for the multilevel logistic regression models in primary education.

	M0 (Null)	M1 (Backgrounds)	M2 (School)	M3 (NC+MCC)	M4 (All)
Fixed effects					
Intercept	$0.10(0.03)^{***}$	$0.14(0.08)^{***}$	$0.07(0.09)^{***}$	$0.41(0.11)^{***}$	$0.38(0.17)^{***}$
ESCS_N1		$0.59(0.03)^{***}$			$0.67(0.03)^{***}$
SCCHANGE_N1		$2.10(0.04)^{***}$			$1.89(0.04)^{***}$
IMMIG_N1		$0.62(0.05)^{***}$			$0.58(0.05)^{***}$
ICTRES_N1		$0.90(0.03)^{***}$			$0.90(0.03)^{***}$
HEDRES_N1		$0.89(0.02)^{***}$			$0.93(0.03)^{***}$
DURECEC_N1		$0.69(0.05)^{***}$			$0.77(0.05)^{***}$
$SC048Q03NA_N2$		$1.01(0.00)^{***}$			$1.01(0.00)^{***}$
BMMJ1_N1		$0.99(0.00)^{***}$			$1.00(0.00)^{***}$
MMINS_N1			$1.00(0.00)^{***}$		$1.00(0.00)^{***}$
IN_NINS_N1			$1.00(0.00)^{***}$		$1.00(0.00)^{***}$
INFOCAR_N1			$0.72(0.02)^{***}$		$0.79(0.02)^{***}$
IN_NINS_N1			$1.00(0.00)^{***}$		$1.00(0.00)^{***}$
DIRINS_N1			$1.30(0.03)^{***}$		$1.24(0.02)^{***}$
TMINS_N1			$1.00(0.00)^{***}$		$1.00(0.00)^{***}$
PERFEED_N1			$1.25(0.03)^{***}$		$1.22(0.02)^{***}$
DISCLIMA_N1			$0.83(0.02)^{***}$		$0.89(0.02)^{***}$

BEINGBULLIED_ N1			$1.30(0.02)^{***}$		$1.13(0.02)^{***}$
BSMJ_N1				$0.97(0.00)^{***}$	$0.98(0.00)^{***}$
PISADIFF_N1				$1.31(0.02)^{***}$	$1.19(0.02)^{***}$
METASUM_N1				$0.69(0.02)^{***}$	$0.75(0.02)^{***}$
STUBMI_N1				$1.01(0.00)^{***}$	$1.00(0.00)^{***}$
GCSELFEFF_N1				$0.87(0.02)^{***}$	$0.89(0.02)^{***}$
BELONG_N1				$0.87(0.02)^{***}$	$0.93(0.02)^{***}$
SCREADDIFF_N1				$1.89(0.02)^{***}$	$1.08(0.02)^{***}$
GCAWARE_N1				0.99(0.02)	$1.01(0.03)^{***}$
COGFLEX_N1				$1.08(0.02)^{***}$	$1.06(0.02)^{***}$
Random effects					
Variance	$0.52(0.05)^{***}$	$0.20(0.03)^{***}$	$0.42(0.04)^{***}$	$0.39(0.04)^{***}$	$0.20(0.07)^{***}$
PEV (%)		62.38	18.62	25.53	0.61
Fit indices					
AIC	22198.69	19804.01	21053.05	20120.57	18072.87
BIC	22215.58	19888.48	21145.96	20213.48	18309.368
Loglik	-11097.34	-9892.01	-10515.53	-10049.29	-9008.43
Deviance	20724.84	19026.15	19749.16	18906.55	17292.51
Chisq		2410.72	1163.65	2096.13	1767.1
$\Pr(> Chisq)$		0.00	0.00	0.00	0.00

Note: p < .05, p < .01, p < .01, p < .001. The reference category to interpret the odds ratio is the value 1 in all cases.



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Lastly, Table 6 presents the regression models for grade repetition in secondary education, with N1 or N2 after the variable name to identify the level to which it belongs. Model 1 explains approximately 51% of the criterion and it is not significantly affected by home ICT resources (ICTRES), with the main risk factor being the number of changes in educational biography (CHANGE). Model 2 explains 32% of the variance; all the variables are statistically significant, although four of them have an odds ratio that is very close to 1. The greatest effect corresponds to student behaviour hindering learning (STUBEHA), a level 2 variable. Model 3, containing NC+MCC predictors, explains 14%. Student resilience (RESILIENCE) is not significant and there are two other predictors with an odds ratio that is very close to 1, with the greatest risk factor found in the variable perception of difficulty of the PISA test (PISADIFF). In the global model, *learning time for sciences* (SMINS), *learning time for mathematics* (MMINS) and *resilience* (RESILIENCE) are not significant and there are another six predictors with an odds ratio that is very close to 1. This set of predictors explains almost 56% of grade repetition in secondary education. The variables with the greatest effect are *number of school changes* (SCCHANGE), *student behaviour hindering learning* (STUBEHA) and *teacher-directed instruction* (DIRINS).

Lastly, it should be noted that model 4 provides the best fit, with a significant reduction in the variance, equivalent to an R^2 of 10.298% compared to model 1, although the reduction in model 1 is also significant in comparison to model 0 (7%), and in models 2 and 3 compared to the null model (5% and 10%, respectively).

TABLE 6. Odds ratio and standard errors for the multilevel logistic regression models in secondary education.

	M0 (Null)	M1 (Backgrounds)	M2 (School)	M3 (NC+MCC)	M4 (All)
Fixed effects					
Intercept	0.21(0.03)***	0.27(0.06)***	0.14(0.07)***	1.00(0.10)	0.70(0.13)***
ESCS_N1		0.67(0.02)***			0.76(0.02)***
CHANGE_N1		2.44(0.03)***			2.24(0.03)***
BMMJ1_N1		0.98(0.01)***			0.99(0.00)***
HEDRES_N1		0.85(0.02)***			0.92(0.02)***
ICTRES_N1		0.91(0.02)			0.91(0.02)***
DURECEC_N1		0.69(0.05)***			0.76(0.04)***
SMINS_N1			1.00(0.00)***		1.00(0.00)***
LMINS_N1			$1.01(0.00)^{***}$		1.00(0.00)***



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MMINS_N1			1.00(0.00)***		1.00(0.00)
					1.00(0.00)
INFOCAR_N1			0.72(0.02)***		0.76(0.02)***
STUBEHA_N2			1.42(0.03)***		1.25(0.02)***
HOMESCH_N1			0.91(0.02)***		0.93(0.02)***
TMINS_N1			1.00(0.00)***		1(0.00)***
ICTOUTSIDE_N	1		0.89(0.02)***		0.96(0.02)***
ENTUSE_N1			$1.10(0.02)^{***}$		$1.13(0.02)^{***}$
INFOJOB2_N1			$1.13(0.02)^{***}$		$1.11(0.02)^{***}$
PERFEED_N1			1.16(0.02)***		$1.15(0.02)^{***}$
INFOJOB1_N1			0.84(0.02)***		0.85(0.02)***
DISCRIM_N1			1.39(0.02)***		1.20(0.02)***
DIRINS_N1			1.21(0.02)***		1.21(0.02)***
BSMJ_N1				0.97(0.00)***	0.98(0.00)***
MASTGOAL_N1				0.78(0.02)***	0.82(0.02)***
PISADIFF_N1				1.23(0.02)***	1.14(0.02)***
GCSELFEFF_N	1			0.86(0.02)***	0.88(0.02)***
RESILIENCE_N	1			1.00(0.02)	0.99(0.02)
METASUM_N1				0.72(0.02)***	0.79(0.02)***
STUBMI_N1				1.01(0.00)***	1.01(0.00)***
GCAWARE_N1				1.08(0.02)***	0.76(0.02)***
Random effects					
Variance	0.70(0.04)***	0.34(0.05)***	0.47(0.04)***	0.60(0.05)***	0.32(0.04)***
PEV (%)		51	32.43	14	55.69
Fit indices					
AIC	32841,56	29736,77	30809,14	29611,70	26708,20
BIC	32858.46	29804.34	30944.28	29696.16	26961.58
Loglik	-16418.78	-14860.38	-15388.57	-14795.85	-13324.10
Deviance	30681.45	28238.71	29013.64	27662.13	25330.76
Chisq		3116.8	2060.3	152.57	6215.6
Pr (> Chisq)	0.70(0.04)***	0.34(0.05)***	0.47(0.04)***	0.60(0.05)***	0.32(0.04)***

Note: p < .05, p < .01, p < .01. The reference category to interpret the odds ratio is the value 1 in all cases.



4. Discussion

In the modelling produced, the most significant contextual indices have proved to be mainly from the first level, that is to say, student characteristics predominate over school characteristics. This result coincides with that obtained by Lopez et al. (2023) in their systematic review of research conducted with Spanish students. Also with other secondary analyses in which the object of study is performance, as in the study by Choi et al. (2018) with data on Spanish students from PISA 2012 and PIRLS 2006 (Progress in International Reading Literacy Study), or by Lopes et al. (2022) with data on Portuguese students from PIRLS 2016.

It can be concluded that the three variables that are most closely linked to grade repetition throughout compulsory education are social status, from the student background group; *learning time*, from the group of constructs associated with school; and the level of expectation, from the NC+MCC set, in a different order in the two stages. In this last group, the results are in line with those found by Asensio et al. (2018) and López et al. (2018), who, without differentiating the stages, establish expectations as one of the most important variables as regards performance. In terms of the importance of the student background variable, the present study is in line with the findings of Carabaña (2013), Choi et al. (2018) or García-Pérez et al. (2014). However, here we provide specific information relating to the differential weight depending on the stage, which may be interpreted as an indicator that the educational system is fulfilling its compensatory function, as

ESCS moves into third place in secondary education, behind a psychological variable and one related to the curriculum. Without a doubt, this evidence represents an interesting contribution by the present study. In the two stages, other significant psychoeducational elements emerge such as, from the set of school-related variables, academic/professional information (ranked tenth in primary education and seventh in secondary education), teacher-directed instruction (thirteenth place in primary education and twenty-ninth in secondary education), perceived teacher feedback and the *climate in the classroom*; and, from the NC+MCC group, perception of the difficulty of the test.

Specifically of note, in primary education only, is the relationship between grade repetition with *bullying* and *flexibility* or *cognitive adaptability* and in secondary education, with *use of ICT*, *learning goals* and *awareness of global issues*. Thus, this stage shows an increase in the number of school-related variables considered important, with a decrease in those related to background, on which the school has less power to act directly, thereby confirming and expanding the findings by Arroyo et al. (2019) with data from PISA 2015.

The emergence of differential models in both stages leads to the conclusion that how grade repetition is treated at the practical and political level needs to take this aspect into account. The existence of clear links between background variables and grade repetition also has implications for practice, as they are a warning of the need to persevere with compensatory ap-

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proaches that improve equity, particularly in primary education. Furthermore, the importance of the time spent on learning indicates practices aimed at efficient management of this variable, without being restricted to simply increasing study time in disadvantaged classes, which is a problem raised, for example, by the policies of summer courses to recover "educational debts" (Battistin & Schizzerotto, 2019). The most equitable solution would be to support the principles of individualisation, early intervention (Choi et al., 2018) and inclusion (Tapia & Álvarez, 2022), in which teacher training and collaborative teaching or team teaching are portrayed as promising strategies in the most successful educational systems (Niemi, 2015). In terms of the level of expectation, it is important to work on motivation, at both the primary and secondary stages (Rhodes et al., 2018; Rodríguez-Rodríguez, 2022). Other specific aspects that should be developed in order to lessen the rate of grade repetition in both stages are feedback on learning, the classroom climate or academic/professional guidance.

The availability of such an extensive database as PISA has enabled the use of machine learning algorithms in this work, which is one of its main contributions. The multilevel study conducted was based on a prior selection that helped to minimise specification error, which is difficult to avoid in educational research. The use of machine learning techniques complements and increases the sensitivity of hierarchical regression models, by providing empirical criteria to maintain the important predictors and dismiss the ones that are irrelevant. Although it has not proved to be so efficient as regards the latter function due to the existence of significant odds ratios, albeit equal or very close to 1, this is a methodology that allows for a quantitative calculation not only of each predictor's worth but also, fundamentally, the value of the model's predictive validity.

In terms of limitations, it should be noted that machine learning methods may overfit models to the current sample's characteristics. Additionally, the PISA background questionnaires are based on self-reporting, which is not free of measurement errors. Moreover, it is worth mentioning that the relationships identified do not imply causality, so for this reason the results of this type of exploratory study should be complemented by confirmatory research posing hypotheses on the direct and indirect effects that apply to the factors involved. In this sense, any educational measure that is adopted to reduce grade repetition, given the relationship of this variable with performance, will have an indirect effect on it, although this aspect should be studied specifically using covariance structure modelling, or designs that are experimental or quasiexperimental. For the future, it would also be advisable to examine the replicability of the results in depth, by comparing the results in participant subgroups, for example, by gender. Likewise, it would be interesting to explore the factors associated with grade repetition by extending the hierarchy levels to include teachers (three-level model) and even autonomous regions (four-level model).



Authors' contributions

Alexander Constante-Amores: Data curation; Formal analysis; Methodology; Writing (original draft).

Delia Arroyo-Resino: Conceptualisation; Formal analysis; Writing (original draft); Writing (review and editing).

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